Econ 422 – Lecture Notes Part I

(These notes are slightly modified versions of lecture notes provided by Stock and Watson, 2007. They are for instructional purposes only and are not to be distributed outside of the classroom.)

The General IV Regression Model

- So far we have considered IV regression with a single endogenous regressor (X) and a single instrument (Z).
- We need to extend this to:
 - o multiple endogenous regressors $(X_1,...,X_k)$
 - \circ multiple included exogenous variables $(W_1, ..., W_r)$

These need to be included for the usual OV reason

- \circ multiple instrumental variables $(Z_1,...,Z_m)$
 - More (relevant) instruments can produce a smaller variance of TSLS: the R^2 of the first stage increases, so you have more variation in \hat{X} .
- Terminology: identification & overidentification

Identification

- In general, a parameter is said to be *identified* if different values of the parameter would produce different distributions of the data.
- In IV regression, whether the coefficients are identified depends on the relation between the number of instruments (*m*) and the number of endogenous regressors (*k*)
- Intuitively, if there are fewer instruments than endogenous regressors, we can't estimate $\beta_1, ..., \beta_k$
 - \circ For example, suppose k = 1 but m = 0 (no instruments)!

Identification, ctd.

The coefficients β_1, \ldots, β_k are said to be:

• exactly identified if m = k.

There are just enough instruments to estimate $\beta_1, ..., \beta_k$.

• *overidentified* if m > k.

There are more than enough instruments to estimate $\beta_1,...,\beta_k$. If so, you can test whether the instruments are valid (a test of the "overidentifying restrictions") – we'll return to this later

• underidentified if m < k.

There are too few instruments to estimate $\beta_1,...,\beta_k$. If so, you need to get more instruments!

The general IV regression model: Summary of jargon

$$Y_i = \beta_0 + \beta_1 X_{1i} + \ldots + \beta_k X_{ki} + \beta_{k+1} W_{1i} + \ldots + \beta_{k+r} W_{ri} + u_i$$

- Y_i is the *dependent variable*
- $X_{1i},...,X_{ki}$ are the *endogenous regressors* (potentially correlated with u_i)
- $W_{1i},...,W_{ri}$ are the *included exogenous variables* or *included exogenous regressors* (uncorrelated with u_i)
- $\beta_0, \beta_1, ..., \beta_{k+r}$ are the unknown regression coefficients
- $Z_{1i},...,Z_{mi}$ are the *m* instrumental variables (the excluded exogenous variables)
- The coefficients are *overidentified* if m > k; *exactly identified* if m = k; and *underidentified* if m < k.

TSLS with a single endogenous regressor

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 W_{1i} + \dots + \beta_{1+r} W_{ri} + u_i$$

- *m* instruments: $Z_{1i},...,Z_m$
- First stage
 - \circ Regress X_1 on *all* the exogenous regressors: regress X_1 on $W_1,...,W_r,Z_1,...,Z_m$ by OLS
 - \circ Compute predicted values \hat{X}_{1i} , i = 1,...,n
- Second stage
 - \circ Regress Y on $\hat{X}_1, W_1, ..., W_r$ by OLS
 - The coefficients from this second stage regression are the TSLS estimators, but *SE*s are wrong
- To get correct SEs, do this in a single step

Example: Demand for cigarettes

$$\ln(Q_i^{cigarettes}) = \beta_0 + \beta_1 \ln(P_i^{cigarettes}) + \beta_2 \ln(Income_i) + u_i$$

 Z_{1i} = general sales tax_i

 Z_{2i} = cigarette-specific tax_i

- Endogenous variable: $ln(P_i^{cigarettes})$ ("one X")
- Included exogenous variable: $ln(Income_i)$ ("one W")
- Instruments (excluded endogenous variables): general sales tax, cigarette-specific tax ("two Zs")
- Is the demand elasticity β_1 overidentified, exactly identified, or underidentified?

Example: Cigarette demand, one instrument

```
. ivreg lpackpc lperinc (lravgprs = rtaxso) if year==1995, r;
IV (2SLS) regression with robust standard errors
                                               Number of obs = 48
                                               F(2, 45) = 8.19
                                               Prob > F = 0.0009
                                               R-squared = 0.4189
                                               Root MSE = .18957
                        Robust
    lpackpc | Coef. Std. Err. t P>|t| [95% Conf. Interval]
   lravgprs | -1.143375 .3723025 -3.07 0.004 -1.893231 -.3935191
    lperinc | .214515 .3117467 0.69 0.495 -.413375 .842405
     _cons | 9.430658 1.259392 7.49 0.000 6.894112 11.9672
Instrumented: lravgprs
            lperinc rtaxso
Instruments:
                               STATA lists ALL the exogenous regressors
                               as instruments - slightly different
                               terminology than we have been using
```

- Running IV as a single command yields correct SEs
- Use, r for heteroskedasticity-robust SEs

Example: Cigarette demand, two instruments

```
m{W} \qquad \qquad m{X} \qquad \qquad m{Z}_1 \qquad \quad m{Z}_2
. ivreg lpackpc lperinc (lravgprs = rtaxso rtax) if year==1995, r;
                                                  Number of obs = 48
IV (2SLS) regression with robust standard errors
                                                  F(2, 45) = 16.17
                                                  Prob > F = 0.0000
                                                  R-squared = 0.4294
                                                  Root MSE = .18786
                         Robust
    lpackpc | Coef. Std. Err. t P>|t| [95% Conf. Interval]
   lravgprs | -1.277424 .2496099 -5.12 0.000 -1.780164 -.7746837
    lperinc | .2804045 .2538894 1.10 0.275 -.230955 .7917641
      _cons | 9.894955 .9592169 10.32 0.000 7.962993 11.82692
Instrumented: lravgprs
Instruments:
             lperinc rtaxso rtax STATA lists ALL the exogenous regressors
                                  as "instruments" - slightly different
                                  terminology than we have been using
```

TSLS estimates,
$$Z = \text{sales tax } (m = 1)$$

$$\ln(Q_i^{cigarettes}) = 9.43 - 1.14 \ln(P_i^{cigarettes}) + 0.21 \ln(Income_i)$$
(1.26) (0.37) (0.31)

TSLS estimates, Z = sales tax, cig-only tax (m = 2)

$$\ln(Q_i^{cigarettes}) = 9.89 - 1.28 \ln(P_i^{cigarettes}) + 0.28 \ln(Income_i)$$
(0.96) (0.25)

- Smaller SEs for m = 2. Using 2 instruments gives more information more "as-if random variation".
- Low income elasticity (not a luxury good); income elasticity not statistically significantly different from 0
- Surprisingly high price elasticity

The General Instrument Validity Assumptions

$$Y_i = \beta_0 + \beta_1 X_{1i} + \ldots + \beta_k X_{ki} + \beta_{k+1} W_{1i} + \ldots + \beta_{k+r} W_{ri} + u_i$$

- (1) Instrument exogeneity: $corr(Z_{1i},u_i) = 0,..., corr(Z_{mi},u_i) = 0$
- (2) *Instrument relevance*: *General case, multiple X's*Suppose the second stage regression could be run using the predicted values from the *population* first stage regression. Then: there is no perfect multicollinearity in this (infeasible) second stage regression.
 - Multicollinearity interpretation...
 - Special case of one X: the general assumption is equivalent to (a) at least one instrument must enter the population counterpart of the first stage regression, and (b) the W's are not perfectly multicollinear.

The IV Regression Assumptions

$$Y_i = \beta_0 + \beta_1 X_{1i} + \ldots + \beta_k X_{ki} + \beta_{k+1} W_{1i} + \ldots + \beta_{k+r} W_{ri} + u_i$$

- 1. $E(u_i|W_{1i},...,W_{ri})=0$
 - #1 says "the exogenous regressors are exogenous."
- 2. $(Y_i, X_{1i}, ..., X_{ki}, W_{1i}, ..., W_{ri}, Z_{1i}, ..., Z_{mi})$ are i.i.d.
 - #2 is not new
- 3. The X's, W's, Z's, and Y have nonzero, finite 4^{th} moments
 - #3 is not new
- 4. The instruments $(Z_{1i},...,Z_{mi})$ are valid.
 - We have discussed this
- Under 1-4, TSLS and its *t*-statistic are normally distributed
- The critical requirement is that the instruments be valid...

Checking Instrument Validity

Recall the two requirements for valid instruments:

- Relevance (special case of one X)
 At least one instrument must enter the population counterpart of the first stage regression.
- 2. Exogeneity

All the instruments must be uncorrelated with the error term: $corr(Z_{1i}, u_i) = 0, ..., corr(Z_{mi}, u_i) = 0$

What happens if one of these requirements isn't satisfied? How can you check? What do you do? If you have multiple instruments, which should you use?

Checking Assumption #1: Instrument Relevance

We will focus on a single included endogenous regressor:

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 W_{1i} + \dots + \beta_{1+r} W_{ri} + u_i$$

First stage regression:

$$X_i = \pi_0 + \pi_1 Z_{1i} + \ldots + \pi_m Z_{mi} + \pi_{m+1} W_{1i} + \ldots + \pi_{m+k} W_{ki} + u_i$$

- The instruments are relevant if at least one of $\pi_1, ..., \pi_m$ are nonzero.
- The instruments are said to be *weak* if all the $\pi_1, ..., \pi_m$ are either zero or nearly zero.
- Weak instruments explain very little of the variation in X, beyond that explained by the W's

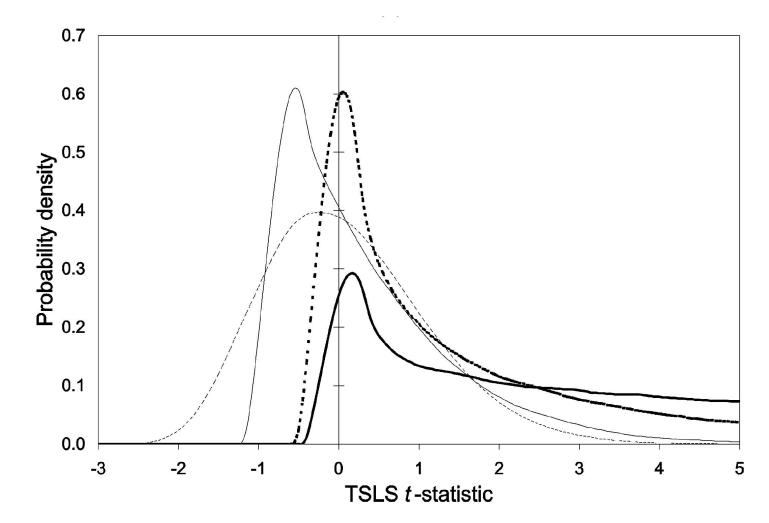
What are the consequences of weak instruments?

If instruments are weak, the sampling distribution of TSLS and its *t*-statistic are not (at all) normal, even with *n* large. Consider the simplest case:

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$
$$X_i = \pi_0 + \pi_1 Z_i + u_i$$

- The IV estimator is $\hat{\beta}_1^{TSLS} = \frac{S_{YZ}}{S_{XZ}}$
- If cov(X,Z) is zero or small, then s_{XZ} will be small: With weak instruments, the denominator is nearly zero.
- If so, the sampling distribution of $\hat{\beta}_1^{TSLS}$ (and its *t*-statistic) is not well approximated by its large-*n* normal approximation...

An example: the sampling distribution of the TSLS t-statistic with weak instruments



Dark line = irrelevant instruments

Dashed light line = strong instruments

Why does our trusty normal approximation fail us?

$$\hat{\beta}_1^{TSLS} = \frac{S_{YZ}}{S_{XZ}}$$

- If cov(X,Z) is small, small changes in s_{XZ} (from one sample to the next) can induce big changes in $\hat{\beta}_1^{TSLS}$
- Suppose in one sample you calculate $s_{XZ} = .00001...$
- Thus the large-*n* normal approximation is a poor approximation to the sampling distribution of $\hat{\beta}_1^{TSLS}$
- A better approximation is that $\hat{\beta}_1^{TSLS}$ is distributed as the *ratio* of two correlated normal random variables (see SW App. 12.4)
- If instruments are weak, the usual methods of inference are unreliable potentially very unreliable.

Measuring the strength of instruments in practice: The first-stage F-statistic

- The first stage regression (one X): Regress X on $Z_1,...,Z_m,W_1,....,W_k$.
- Totally irrelevant instruments \Leftrightarrow *all* the coefficients on Z_1, \ldots, Z_m are zero.
- The *first-stage F-statistic* tests the hypothesis that $Z_1,...,Z_m$ do not enter the first stage regression.
- \bullet Weak instruments imply a small first stage F-statistic.

Checking for weak instruments with a single X

- Compute the first-stage F-statistic.

 Rule-of-thumb: If the first stage F-statistic is less than

 10, then the set of instruments is weak.
- If so, the TSLS estimator will become more severely biased, and statistical inferences (standard errors, hypothesis tests, confidence intervals) can be misleading.
- Note that simply rejecting the null hypothesis that the coefficients on the Z's are zero isn't enough you actually need substantial predictive content for the normal approximation to be a good one.
- There are more sophisticated things to do than just compare *F* to 10 but they are beyond this course.

What to do if you have weak instruments?

- Get better instruments (!)
- If you have a number of instruments and perhaps some are much weaker than others; then, one approach is to drop some of the weaker ones (dropping an irrelevant instrument will increase the first-stage *F*). This leads to a possible pretesting problem, however.
- If you only have a few instruments, and all are weak, then you need to do some IV analysis other than TSLS...
 - \circ Separate the problem of estimation of β_1 and construction of confidence intervals
 - This seems odd, but if TSLS isn't normally distributed, it makes sense (right?)

Estimation with weak instruments

- TSLS estimator become more severely biased if instruments are weak or irrelevant.
- However, some estimators have a distribution more centered around β_1 than does TSLS
- One such estimator is the limited information maximum likelihood estimator (LIML)
- The LIML estimator
 - o can be derived as a maximum likelihood estimator
 - o is the value of β_1 that minimizes the *p*-value of the AR test(!)

Checking Assumption #2: Instrument Exogeneity

- Instrument exogeneity: *All* the instruments are uncorrelated with the error term: $corr(Z_{1i},u_i) = 0,...,$ $corr(Z_{mi},u_i) = 0$
- If the instruments are correlated with the error term, the first stage of TSLS doesn't successfully isolate a component of X that is uncorrelated with the error term, so \hat{X} is correlated with u and TSLS is inconsistent.
- If there are more instruments than endogenous regressors, it is possible to test *partially* for instrument exogeneity.

Testing overidentifying restrictions

Consider the simplest case:

$$Y_i = \beta_0 + \beta_1 X_i + u_i,$$

- Suppose there are two valid instruments: Z_{1i} , Z_{2i}
- Then you could compute two separate TSLS estimates.
- Intuitively, if these 2 TSLS estimates are very different from each other, then something must be wrong: one or the other (or both) of the instruments must be invalid.
- The *J*-test of overidentifying restrictions makes this comparison in a statistically precise way.
- This can only be done if #Z's > #X's (overidentified).

Suppose #instruments = m > # X's = k (overidentified) $Y_i = \beta_0 + \beta_1 X_{1i} + ... + \beta_k X_{ki} + \beta_{k+1} W_{1i} + ... + \beta_{k+r} W_{ri} + u_i$

The *J*-test of overidentifying restrictions

The *J*-test can be constructed as follows:

- 1. First estimate the equation of interest using TSLS and all m instruments; compute the predicted values \hat{Y}_i , using the actual X's (not the \hat{X} 's used to estimate the second stage)
- 2. Compute the residuals $\hat{u}_i = Y_i \hat{Y}_i$
- 3. Regress \hat{u}_i against $Z_{1i},...,Z_{mi}, W_{1i},...,W_{ri}$
- 4. Compute the F-statistic testing the hypothesis that the coefficients on $Z_{1i},...,Z_{mi}$ are all zero;
- 5. The *J-statistic* is J = mF

J = mF, where F = the F-statistic testing the coefficients on $Z_{1i},...,Z_{mi}$ in a regression of the TSLS residuals against $Z_{1i},...,Z_{mi}, W_{1i},...,W_{ri}$.

Distribution of the *J*-statistic

- Under the null hypothesis that all the instruments are exogeneous, *J* has a chi-squared distribution with *m*–*k* degrees of freedom
- If m = k, J = 0 (does this make sense?)
- If some instruments are exogenous and others are endogenous, the *J* statistic will be large, and the null hypothesis that all instruments are exogenous will be rejected.

Checking Instrument Validity: Summary

The two requirements for valid instruments:

- **1.** *Relevance* (special case of one X)
 - At least one instrument must enter the population counterpart of the first stage regression.
 - If instruments are weak, then the TSLS estimator will be more severely biased and the *t*-statistic has a non-normal distribution
 - To check for weak instruments with a single included endogenous regressor, check the first-stage *F*
 - \circ If F>10, instruments are strong use TSLS
 - \circ If F<10, weak instruments take some action

2. Exogeneity

- *All* the instruments must be uncorrelated with the error term: $corr(Z_{1i}, u_i) = 0, ..., corr(Z_{mi}, u_i) = 0$
- We can partially test for exogeneity: if m>1, we can test the hypothesis that all are exogenous, against the alternative that as many as m-1 are endogenous (correlated with u)
- The test is the *J*-test, constructed using the TSLS residuals.